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**Statistical interpretation of data — Tests for  
departure from the normal distribution**

*Interprétation statistique des données — Tests pour les écarts à la  
distribution normale*

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## Foreword

ISO (the International Organization for Standardization) is a worldwide federation of national standards bodies (ISO member bodies). The work of preparing International Standards is normally carried out through ISO technical committees. Each member body interested in a subject for which a technical committee has been established has the right to be represented on that committee. International organizations, governmental and non-governmental, in liaison with ISO, also take part in the work. ISO collaborates closely with the International Electrotechnical Commission (IEC) on all matters of electrotechnical standardization.

Draft International Standards adopted by the technical committees are circulated to the member bodies for voting. Publication as an International Standard requires approval by at least 75 % of the member bodies casting a vote.

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International Standard ISO 5479 was prepared by Technical Committee ISO/TC 69, *Applications of statistical methods*, Subcommittee SC 6, *Measurement methods and results*.

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Annexes A and B of this International Standard are for information only.

## Introduction

Many of the statistical methods recommended in International Standards, such as those described in ISO 2854<sup>[1]</sup>, are based on the assumption that the random variable(s) to which these methods apply are independently distributed according to a normal distribution with one or both of its parameters unknown.

The following question therefore arises. Is the distribution that is represented by the sample sufficiently close to the normal distribution that the methods provided by these International Standards can be used reliably?

There is no simple yes or no answer to this question which is valid in all cases. For this reason a large number of "tests of normality" have been developed, each of which is more or less sensitive to a particular feature of the distribution under consideration; e.g. asymmetry or kurtosis.

Generally the test used is designed to correspond to a predetermined *a priori* risk that the hypothesis of normality is rejected even if it is true (error of the first kind). On the other hand, the probability that this hypothesis is not rejected when it is not true (error of the second kind) cannot be determined unless the alternative hypothesis (i.e. that which is opposed to the hypothesis of normality) can be precisely defined. This is not possible in general and, furthermore, it requires computational effort. For a distinct test, this risk is particularly large if the sample size is small.

# Statistical interpretation of data — Tests for departure from the normal distribution

## 1 Scope

**1.1** This International Standard gives guidance on methods and tests for use in deciding whether or not the hypothesis of a normal distribution should be rejected, assuming that the observations are independent.

**1.2** Whenever there are doubts as to whether the observations are normally distributed, the use of a test for departure from the normal distribution may be useful or even necessary. In the case of robust methods, however (i.e. where the results are only altered very slightly when the real probability distribution of the observations is not a normal distribution), a test for departure from the normal distribution is not very helpful. This is the case, for example, when the mean of a single random sample of observations is to be checked against a given theoretical value using a *t*-test.

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**1.3** It is not strictly necessary to use such a test whenever one refers to statistical methods based on the hypothesis of normality. It is possible that there is no doubt at all as to the normal distribution of the observations, whether theoretical (e.g. physical) reasons are present which confirm the hypothesis or because this hypothesis is deemed to be acceptable according to prior information.

**1.4** The tests for departure from the normal distribution selected in this International Standard are primarily intended for complete data, not grouped data. They are unsuitable for censored data.

**1.5** The tests for departure from the normal distribution selected in this International Standard may be applied either to observed values or to functions of them, such as the logarithm or the square root.

**1.6** Tests for departure from the normal distribution are very ineffective for samples of size less than eight. Accordingly, this International Standard is restricted to samples of eight or more.

## 2 Normative reference

The following standard contains provisions which, through reference in this text, constitute provisions of this International Standard. At the time of publication, the edition indicated was valid. All standards are subject to revision, and parties to agreements based on this International Standard are encouraged to investigate the possibility of applying the most recent edition of the standard indicated below. Members of IEC and ISO maintain registers of currently valid International Standards.

ISO 3534-1:1993, *Statistics — Vocabulary and symbols — Part 1: Probability and general statistical terms.*

### 3 Definitions and symbols

#### 3.1 Definitions

For the purposes of this International Standard, the definitions given in ISO 3534-1 apply.

#### 3.2 Symbols

|              |   |
|--------------|---|
| $a_k$        | coefficient of the Shapiro-Wilk test  |
| $A$          | auxiliary quantity for the Epps-Pulley test   |
| $b_2$        | empirical kurtosis  |
| $\sqrt{b_1}$ | empirical skewness  |
| $B$          | auxiliary quantity for the Epps-Pulley test   |
| $E$          | expectation   |
| $G_j$        | auxiliary quantity for the joint test using several independent samples                   |
| $h$          | number of consecutive samples   |
| $H_0$        | null hypothesis   |
| $H_1$        | alternative hypothesis  |
| $k$          | within the sample, arranged in non-decreasing order, the number of the observed value $x$ |
| $m_j$        | central moment of order $j$ of the sample   |
| $n$          | sample size   |
| $p$          | probability associated with the $p$ -quantile of a distribution                           |
| $P$          | probability   |
| $P_k$        | probability associated with $X_{(k)}$   |
| $S$          | auxiliary quantity for the Shapiro-Wilk test  |
| $T$          | test statistic  |
| $T_{EP}$     | test statistic of the Epps-Pulley test  |
| $u_p$        | $p$ -quantile of the standardized normal distribution                                     |
| $v_j$        | auxiliary quantity for the joint test using several independent samples                   |
| $W$          | test statistic of the Shapiro-Wilk test   |
| $W_j$        | auxiliary quantity for the joint test using several independent samples                   |
| $x$          | value of $X$  |
| $X$          | random variable   |
| $x_{(j)}$    | $j^{\text{th}}$ value in the sample, arranged in non-decreasing order                     |
| $x_{(k)}$    | $k^{\text{th}}$ value in the sample, arranged in non-decreasing order                     |
| $\bar{x}$    | arithmetic average  |
| $\alpha$     | significance level  |

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|                     |   |
|---------------------|---|
| $\beta$             | probability of an error of the second kind                              |
| $\beta_2$           | kurtosis of the population  |
| $\beta_{2-3}$       | excess of the population  |
| $\sqrt{\beta_1}$    | skewness of the population  |
| $\gamma$            | auxiliary quantity for the joint test using several independent samples |
| $\gamma_{(n)}$      | coefficient of the joint test using several independent samples         |
| $\delta$            | auxiliary quantity for the joint test using several independent samples |
| $\delta_{(n)}$      | coefficient of the joint test using several independent samples         |
| $\varepsilon$       | auxiliary quantity for the joint test using several independent samples |
| $\varepsilon_{(n)}$ | coefficient of the joint test using several independent samples         |
| $\mu$               | expectation   |
| $\mu_2$             | variance of the population  |
| $\mu_3$             | central moment of the third order of the population                     |
| $\mu_4$             | central moment of the fourth order of the population                    |
| $\sigma$            | standard deviation of the population ( $=\sqrt{\mu_2}$ )                |

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## 4 General

**4.1** There are several categories of tests for departure from normality. In this International Standard, graphical methods, moment tests, regression tests and characteristic function tests are considered. Chi-squared tests are appropriate for grouped data only but, because grouping results in a loss of information, they are not considered in this International Standard.

**4.2** If no additional information about the sample is available, it is recommended first to do a normal probability plot; i.e. to plot the cumulative distribution function of the observed values on normal probability graph paper consisting of a system of coordinate axes where the cumulative distribution function of the normal distribution is represented by a straight line.

This method, which is described in clause 5, allows one to "see" immediately whether the distribution observed is close to the normal distribution or not. With this additional information it can be decided whether to carry out a directional test, or to carry out either a regression test or a characteristic function test, or no test at all. In addition, although such a graphical representation cannot be considered as a rigorous test, the summary information that it provides is an essential supplement to any test for departure from the normal distribution. In the case of rejection of the null hypothesis it is often possible to envisage by this means the type of alternative that might be applicable.

**4.3** A test for departure from the normal distribution is a test of the null hypothesis that the sample consists of  $n$  independent observations coming from one and the same normal distribution. It consists of the calculation of a function  $T$  of the observations, which is called the test statistic. The null hypothesis of a normal distribution is then not rejected or rejected depending on whether or not the value of  $T$  lies within a set of values near to the expected value that corresponds to the normal distribution.

**4.4** The **critical region** of the test is the set of values of  $T$  that leads to the rejection of the null hypothesis. The **significance level** of the test is the probability  $P$  of obtaining a value of  $T$  within the critical region when the null hypothesis is correct. This level gives the probability of erroneously rejecting the null hypothesis (error of the first kind).

The boundary of the critical region is (or, in the case of a two-sided test, the boundaries of the critical region are) the critical value(s) of the test statistic.

**4.5** The **power** of the test is the probability of rejecting the null hypothesis when it is incorrect. A high power corresponds to a low probability of not rejecting the null hypothesis erroneously (error of the second kind).

It should be emphasized that the power of a test (i.e. for a given situation, the probability that the null hypothesis of a normal distribution will be rejected if it is wrong) increases as the number of observations increases. For example, a departure from the normal distribution which would become apparent when using a test for departure from the normal distribution on a large sample might not be detected by the same test if there were fewer observations.

**4.6** A distinction is made between two categories of tests for departure from the normal distribution. When the form of departure from the normal distribution is specified in the alternative hypothesis, then the test is a **directional test**. However, when the form of departure from the normal distribution is not specified in the alternative hypothesis, the test is an **omnibus test**.

In a directional test, the critical region is determined in such a way that the power of the test reaches its maximum value. In an omnibus test, it is necessary to divide the critical region in such a way that the critical region consists of those values of the test statistic which lie far away from the expected value.

If assumptions are present about the type of departure from the normal distribution, i.e. when a distribution is envisaged whose asymmetry or kurtosis is different from that of the normal distribution, a directional test should be applied, because its power is greater than the power of an omnibus test.

**4.7** Note that a directional test is essentially one-sided. In the case of asymmetry, for example, it centres either on positive asymmetry or on negative asymmetry. However, when several alternatives are considered jointly, the test is multidirectional. This is the case particularly when a non-null asymmetry and a kurtosis different from that of the normal distribution are considered together.

**4.8** Tables 8 to 14 and figure 9 allow the tests to be performed for the most usual levels of  $\alpha$ ; i.e.  $\alpha = 0,05$  and  $\alpha = 0,01$ . The level of significance has to be stipulated before the test is performed. Note that a test may result in the rejection of the null hypothesis at the 0,05 level and the non-rejection of this same hypothesis at the 0,01 level.

**4.9** During computation of test statistics, it is necessary to use at least six significant digits. Subtotals, intermediate results and auxiliary quantities shall not be rounded to less than six significant digits.

## 5 Graphical method

**5.1** The cumulative distribution function of the observed values is plotted on normal probability graph paper. On this paper, one of the axes (in this International Standard it is the vertical axis) is non-linearly scaled according to the area under the standardized normal distribution function and is marked with the corresponding values of the cumulative relative frequency. The other axis is linearly scaled for the ordered values of  $X$ . The cumulative distribution function of the variable  $X$  then approximates to a straight line.

Sometimes these two axes are interchanged with each other. Furthermore, if a normalizing transformation of the variable  $X$  is made, the linear scale may be replaced by a logarithmic, quadratic, reciprocal or other scale.

Figure 1 gives an example of normal probability graph paper. On the vertical axis the values of the cumulative relative frequency are given as percentages, while the horizontal axis has an arbitrary linear scale.



A sheet of blank normal probability graph paper is provided in annex A.

If a plot on this paper gives a set of points that appears to be scattered around a straight line, this provides crude support for the assumption that the sample can reasonably be regarded as having come from a normal distribution.

However, if there is a systematic departure from the straight line, the plot often suggests the type of distribution to be taken into consideration.

The importance of this approach is that it easily provides visual information on the type of departure from the normal distribution.

If the graph indicates that the data come from a shaped distribution (e.g. if the graph of the cumulative distribution function is as shown in figure 5 or 6), a transformation of the data might result in a normal distribution.

If the graph indicates that the data do not come from a simple homogeneous distribution, but rather from a mixture of two or more homogeneous subpopulations (e.g. if the graph of the cumulative distribution function is as shown in figure 7), it is recommended that the subpopulations be identified and the analysis on each subpopulation be continued separately.

It should be kept in mind that such a plot is in no way a test for departure from the normal distribution in the strict sense. In the case of small samples, pronounced curves may occur for normal distributions, whilst for large samples slight curves may indicate non-normal distributions.

**5.2** The graphical procedure consists of arranging the observed values  $(x_{(1)}, x_{(2)}, \dots, x_{(n)})$  in non-decreasing order, and then plotting

$$P_k = (k - 3/8)/(n + 1/4)$$

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... (1)

against  $x_{(k)}$  on normal probability graph paper.

NOTE 1 Commonly used alternatives to equation (1) are [ISO 5479:1997](https://standards.iteh.ai/catalog/standards/sist/2f0bcfbe-9d90-4d03-8a48-5d75243a0b03/iso-5479-1997)  
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$$P_k = (k - 1/2)/n$$

and

$$P_k = k/(n + 1)$$

These are poorer approximations to the normal distribution function of the expected order statistics,  $F[E(X_{(k)})]$ , and their use is not recommended.

**5.3** An example of how normal probability graph paper is used is shown in figure 2.

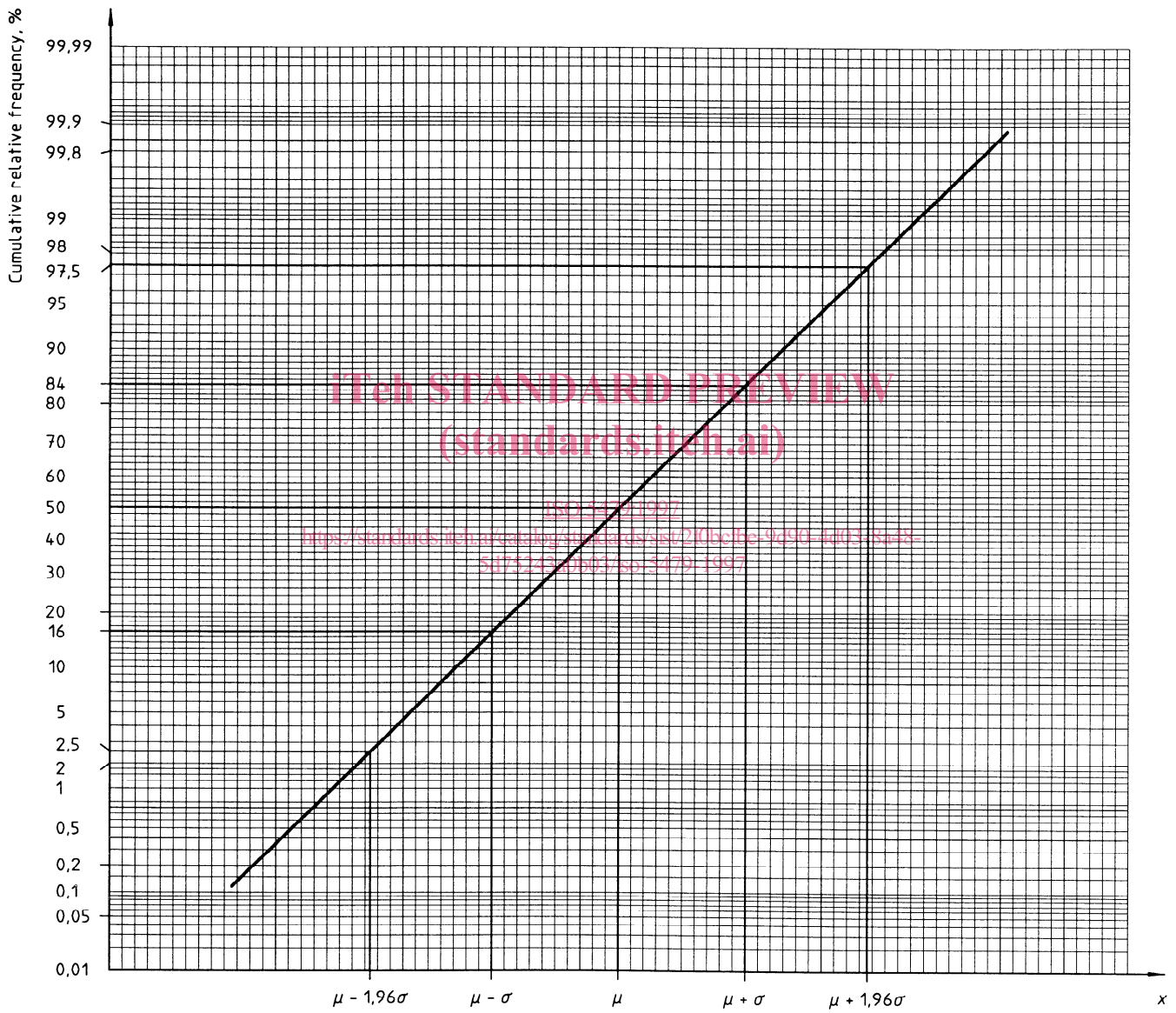


Figure 1 — Annotated normal probability graph paper

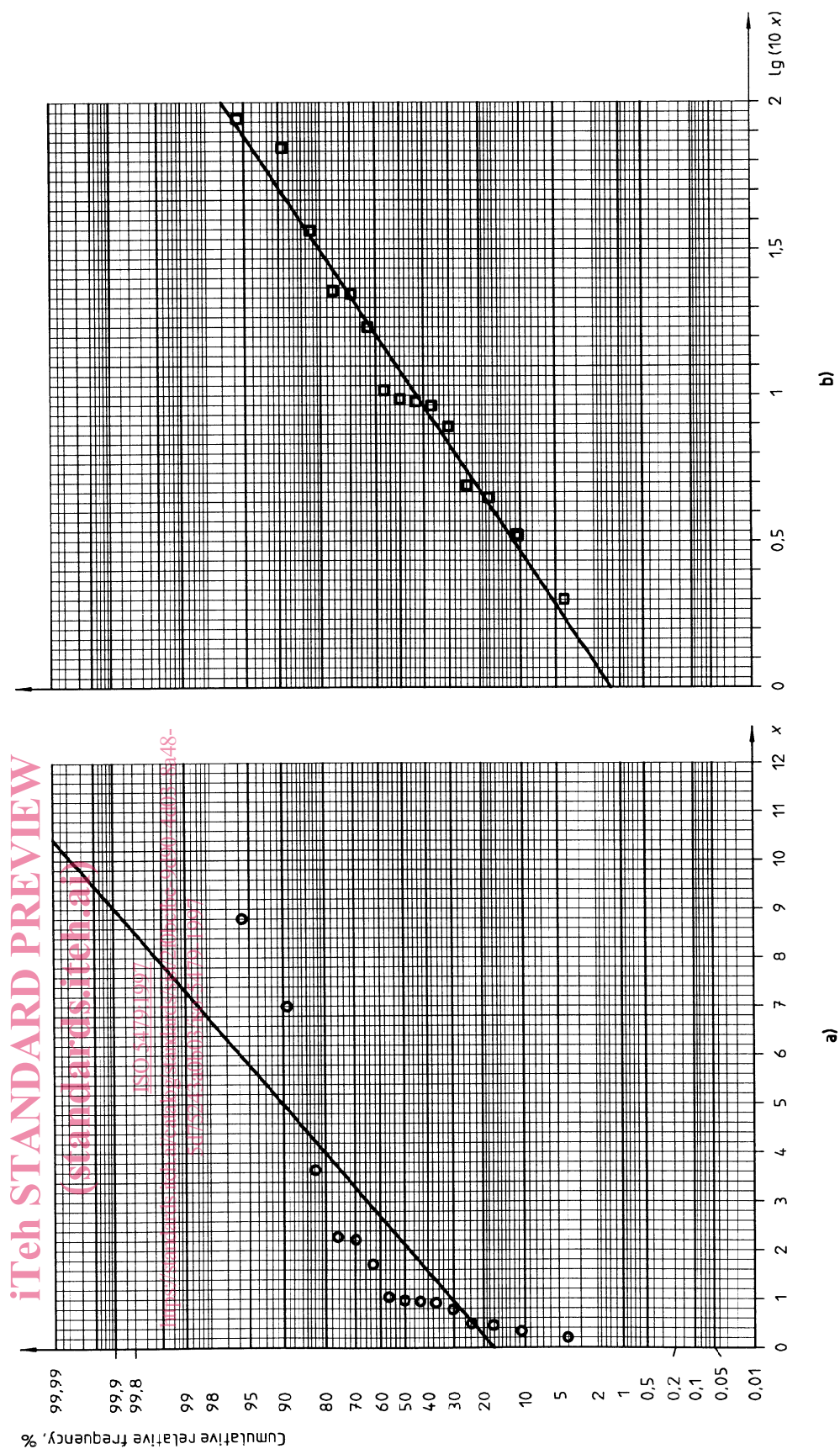


Figure 2 — Graph of a series of observations on normal probability graph paper

Table 1 shows the values  $x_{(k)}$  in non-decreasing order of the result of a series of 15 independent rotating-bend fatigue tests.

**Table 1 – Results,  $x_{(k)}$  of a series of 15 rotating-bend fatigue tests and corresponding values of  $\lg(10 x_{(k)})$**

| $k$ | $P = \frac{k-3/8}{n+1/4}$ | $x_{(k)}$ | $\lg(10x_{(k)})$ |
|-----|---------------------------|-----------|------------------|
| 1   | 0,041                     | 0,200     | 0,301            |
| 2   | 0,107                     | 0,330     | 0,519            |
| 3   | 0,172                     | 0,445     | 0,648            |
| 4   | 0,238                     | 0,490     | 0,690            |
| 5   | 0,303                     | 0,780     | 0,892            |
| 6   | 0,369                     | 0,920     | 0,964            |
| 7   | 0,434                     | 0,950     | 0,978            |
| 8   | 0,500                     | 0,970     | 0,987            |
| 9   | 0,566                     | 1,040     | 1,017            |
| 10  | 0,631                     | 1,710     | 1,233            |
| 11  | 0,697                     | 2,220     | 1,346            |
| 12  | 0,762                     | 2,275     | 1,357            |
| 13  | 0,828                     | 3,650     | 1,562            |
| 14  | 0,893                     | 7,000     | 1,845            |
| 15  | 0,959                     | 8,800     | 1,944            |

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NOTE 2 In table 1 and the following examples, the units for the observations are omitted because they are not relevant for the tests in this International Standard.

By associating the probability

$$P_k = (k - 3/8)/(n + 1/4)$$

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with the  $k$ th smallest  $x_{(k)}$ , the series of points shown in figure 2a) is obtained. It is immediately seen from this graph that these points do not form a straight line. However, if  $x_{(k)}$  is replaced by  $\lg(10 x_{(k)})$ , the new graph [figure 2b)] leads to a series of points which this time lie acceptably close to a straight line.

The hypothesis of a normal distribution of the logarithm of the observations therefore seems adequate.

**5.4** It should be noted that extreme observed values have greater variance than middle values. Therefore, and since the scale for the cumulative relative frequency widens towards the extremes, a few values at either end of the cumulative distribution which distinctly depart from the straight line defined by the middle values cannot be regarded as indicators of departure from the normal distribution.

The larger the sample size, the more reliable are the conclusions that can be derived from the shape of the graph.

If the graph of the cumulative distribution function of the observed values is such that the large values tend to be well below the straight line defined by the other values, a transformation such as

$$y = \log x$$

or

$$y = \sqrt{x}$$

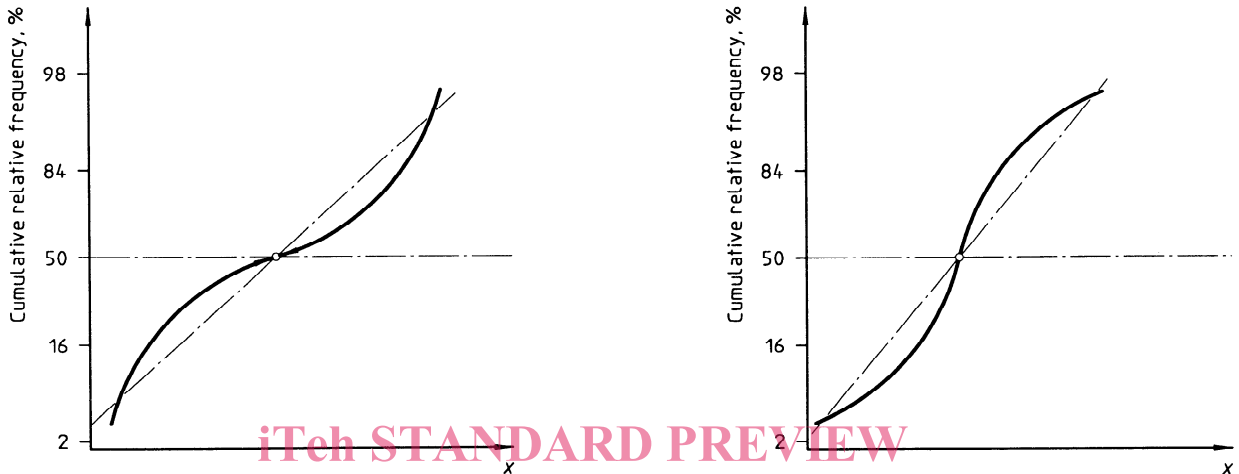
will generally lead to a graph that conforms more to a straight line [see figure 2b) and figure 5].

The upper parts of figures 3 to 7 show the cumulative distribution function in comparison with the corresponding density function shown in the lower part of each figure.

If the graph of the cumulative distribution function of the observed values is as shown in figure 3 or 4, the corresponding frequency distribution is of kurtosis in default (platykurtic) or of kurtosis in excess (leptokurtic), respectively.

The graphs of the cumulative distribution functions shown in figures 5 and 6 correspond to a density function with positive skewness and negative skewness respectively.

Figure 7 shows the cumulative distribution function and the density function of a superposition of two different density functions.



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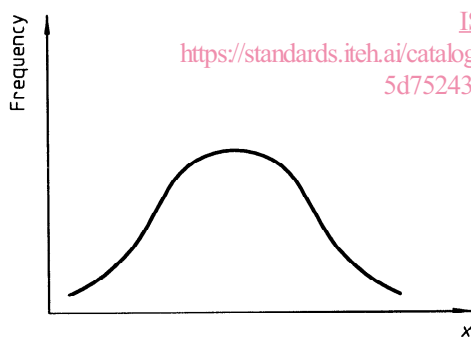


Figure 3 — Density function with kurtosis in default

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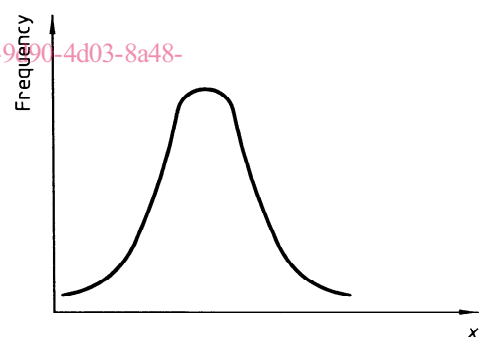


Figure 4 — Density function with kurtosis in excess